

**MVE055 / MVE051 / MSG810 Matematisk statistik och diskret matematik**

Exam 29 August 2018, 14:00 - 18:00

**Allowed aids:** Chalmers-approved calculator  
and one (two-sided) A4 sheet of paper with your own notes.  
Total number of points: 30. To pass, at least 12 points are needed.  
Note: All answers should be motivated.

## Solutions

1. (a) Since  $P(X = k) = \frac{1}{4}, k \in \{-3, -1, 1, 3\}$  then  $P(Y = r) = \frac{1}{2}, r \in \{1, 9\}$ . We can compute  $\mathbb{E}[X] = 0$  and similarly  $\mathbb{E}[X^3] = 0$  since  $X$  is symmetric around 0. We observe now that, since  $Y = X^2, XY = XX^2 = X^3$ , implying that  $\mathbb{E}[XY] = \mathbb{E}[X^3] = 0$ . Thus,

$$\text{Cov}(X, Y) = \mathbb{E}[XY] - \mathbb{E}[X]\mathbb{E}[Y] = \mathbb{E}[X^3] - 0 = 0.$$

Hence,  $X$  and  $Y$  are uncorrelated.

- (b)  $X, Y$  are not independent since, for example,

$$P(X = -1, Y = 9) = 0 \neq P(X = -1)P(Y = 9).$$

2. Chebychev's inequality: Let  $X$  be a random variable such that  $\mathbb{E} = \mu, \text{Var}(X) = \sigma^2$ . If  $0 < \sigma^2 < \infty$ , then for any  $a > 0$  it holds

$$P(|X - \mu| \geq a) \leq \frac{\sigma^2}{a^2}.$$

Proof: Define  $Y = |X - \mu|$ . For any  $a > 0$  we define

$$Z = \begin{cases} a^2 & Y \geq a \\ 0 & \text{otherwise} \end{cases}$$

By definition  $Z \leq Y^2$ , which implies

$$\mathbb{E}[Z] \leq \mathbb{E}[Y^2]. \tag{1}$$

Moreover, we observe

$$\mathbb{E}[Z] = a^2 P(Y \geq a) = a^2 P(|X - \mu| \geq a)$$

and

$$\mathbb{E}[Y^2] = \mathbb{E}[|X - \mu|^2] = \text{Var}(X) = \sigma^2.$$

Replacing the above equations into Equation 1 we obtain

$$a^2 P(|X - \mu| \geq a) \leq \text{Var}(X)$$

which can be rewritten as

$$P(|X - \mu| \geq a) \leq \frac{\text{Var}(X)}{a^2} = \frac{\sigma^2}{a^2}.$$

3. (a) Given that  $P(X \geq x + 1 | X \geq x) = 1 - p$ , for all  $x = 1, 2, 3, \dots$  we have by definition of conditional probability, for all  $x = 1, 2, 3, \dots$ ,

$$1 - p = P(X \geq x + 1 | X \geq x) = \frac{P(X \geq x + 1, X \geq x)}{P(X \geq x)} = \frac{P(X \geq x + 1)}{P(X \geq x)}.$$

- (b) Using the previous point, for all  $x = 1, 2, 3, \dots$ ,

$$P(X \geq x + 1) = (1 - p)P(X \geq x) = (1 - p)^2 P(X \geq x - 1) = \dots = (1 - p)^x P(X \geq 1) = (1 - p)^x,$$

where  $P(X \geq 1) = 1$  since  $X$  is always greater or equal to 1.

- (c)

$$P(X = x) = P(X \geq x) - P(X \geq x + 1) = (1 - p)^x - (1 - p)^{x + 1} = (1 - p)^x (1 - (1 - p)) = p(1 - p)^x.$$

$X$  has a Geometric distribution with parameter  $p$ .

4. (a) TRUE: the minimum value  $U$  can assume is  $a$ . Thus the minimum value  $V$  can assume is  $s + at$ . Similarly, the maximum value of  $V$  is  $s + bt$ . For any value  $v \in [s + at, s + bt]$  the density of  $V$  is

$$f_V(v) = f_U\left(\frac{v - s}{t}\right) \left|\frac{1}{t}\right| = \frac{1}{t(b - a)}$$

which is the density of a uniform random variable on the interval  $[s + at, s + bt]$ .

- (b) FALSE: see previous motivation.

- (c) FALSE: The expected value of  $U$  is  $\mathbb{E}[U] = \frac{b + a}{2}$ . Using the linearity of expectation, we obtain

$$\mathbb{E}[V] = \mathbb{E}[s + tU] = s + t + \mathbb{E}[U] = s + t \frac{b + a}{2}$$

5. The  $100(1 - \alpha)\%$  standard confidence interval for the mean  $\mu$  is given by

$$I_\alpha = \left[ \bar{X} - z_{\frac{\alpha}{2}} \frac{\sigma}{\sqrt{n}}, \bar{X} + z_{\frac{\alpha}{2}} \frac{\sigma}{\sqrt{n}} \right],$$

where  $\bar{X}$  denotes the sample mean and  $z_{\frac{\alpha}{2}}$  the point such that  $P(N(0, 1) \geq z_{\frac{\alpha}{2}}) = \frac{\alpha}{2}$ . The interval  $I_\alpha$  has width

$$\text{width of } I_\alpha = 2z_{\frac{\alpha}{2}} \frac{\sigma}{\sqrt{n}}. \quad (2)$$

- (a) If  $\alpha$  increases, the width of  $I_\alpha$  decreases. In fact, when  $\alpha$  increases, we are requiring a lower probability that the confidence interval will contain the mean  $\mu$ . Thus, the width of  $I_\alpha$  will be smaller.
- (b) If  $\sigma^2$  decreases then the width decreases. In particular, when the variance decreases by a factor 4, the width of the confidence interval considered will halve.
- (c) If  $n$  doubles, the width of the confidence interval considered here will decrease by a factor  $\sqrt{2}$ .
6. (a) The transition matrix  $P$  in canonical form is (rows/columns refers in order to the states 1, 2, 0, 3 in this order)

$$P = \begin{bmatrix} 0 & 0.25 & 0.75 & 0 \\ 0.75 & 0 & 0 & 0.25 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 1 \end{bmatrix} = \begin{bmatrix} Q & R \\ 0 & I \end{bmatrix}$$

- (b) The fundamental matrix is  $N = (I - Q)^{-1} = \begin{bmatrix} \frac{16}{13} & \frac{4}{13} \\ \frac{12}{13} & \frac{1}{13} \end{bmatrix}$ . To obtain the probability of bankrupt, i.e. the probability of being absorbed in the state 0, we need to compute

$$B = NR = \begin{bmatrix} \frac{16}{13} & \frac{4}{13} \\ \frac{12}{13} & \frac{1}{13} \end{bmatrix} \begin{bmatrix} \frac{3}{4} & 0 \\ 0 & \frac{1}{4} \end{bmatrix} = \begin{bmatrix} \frac{12}{13} & \frac{1}{13} \\ \frac{9}{13} & \frac{4}{13} \end{bmatrix}.$$

Since the chain starts in the state 2, the probability of being absorbed in 0 is  $\frac{9}{13} = 0.692$ .